Instrumental Variable II

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Linear Methods in Causal Inference POL1784

Review

- We discussed how to identify causal effects when there exists non-compliance.
- In this case, treatment assignment Z_i no longer equals treatment status D_i.
- We can always identify the intention-to-treat effect but may care about the local average treatment effect (LATE), or the effect on the compliers.
- It can be identified if 1. Z_i is randomly assigned; 2. it affects Y_i only through D_i; 3. it changes the value of D_i monotonically.
- ► Then, we can estimate the LATE using the Wald estimator.

- ► We called treatment assignment Z_i under non-compliance an instrumental variable (IV).
- But the idea of using an IV for causal inference was proposed in a very different context.
- It was introduced by economists to study simultaneous structural equations.
- People used to believe that we can describe a social system (e.g., the US economy) with a large number of equations.
- Variables on the left-hand side are referred to as endogenous, while those only appear on the right-hand side are called exogenous.
- Identification means that we can estimate coefficients in these models consistently.
- Instrumental variables were proposed as a solution to the identification problem.

- ► An economist observes the everyday price and quantity of transaction for fish in a market over N days: (P_i, Q_i)^N_{i=1}.
- She wants to identify the demand curve of fish: $P_i = d(Q_i)$.
- ▶ But we only have the price and quantity at equilibrium, which is also affected by the supply curve: P_i = s(Q_i).
- Suppose both the demand and the supply curves are linear:

$$\begin{aligned} P_{di} &= a_d - b_d Q_{di} + \varepsilon_{di}, \\ P_{si} &= a_s + b_s Q_{si} + \varepsilon_{si}. \end{aligned}$$

We know that at equilibrium,

$$P_{i} = \frac{a_{s}b_{d} + a_{d}b_{s} + b_{s}(\varepsilon_{di} - \varepsilon_{si})}{b_{d} + b_{s}},$$
$$Q_{i} = \frac{a_{d} - a_{s} + \varepsilon_{di} - \varepsilon_{si}}{b_{d} + b_{s}}.$$

Even with a large sample, we can only estimate

$$rac{a_sb_d+a_db_s}{b_d+b_s}$$
 and $rac{a_d-a_s}{b_d+b_s}$

but not (a_d, b_d) .

Suppose there is a shock Z_i on the supply side such as storms at sea:

$$P_{di} = a_d - b_d Q_{di} + \varepsilon_{di},$$

$$P_{si} = a_s + b_s Q_{si} + c_s Z_i + \varepsilon_{si},$$

$$Z_i \perp (\varepsilon_{di}, \varepsilon_{si}).$$

 \triangleright Z_i is known as an instrumental variable.

▶ Now, at equilibrium, we have

$$P_{i} = \frac{a_{s}b_{d} + a_{d}b_{s} + c_{s}b_{d}Z_{i} + b_{s}(\varepsilon_{di} - \varepsilon_{si})}{b_{d} + b_{s}},$$
$$Q_{i} = \frac{a_{d} - a_{s} - c_{s}Z_{i} + \varepsilon_{di} - \varepsilon_{si}}{b_{d} + b_{s}}.$$

 From these relationships, we can obtain estimates of the two slopes

$$rac{c_s b_d}{b_d + b_s}$$
 and $rac{-c_s}{b_d + b_s}.$

Their ratio allows us to identify the parameter b_d.

IV in structural models

- ► Let's consider an economic model of income and education.
- An individual *i* maximizes her return by deciding whether to attend college, D_i ∈ {0,1}.
- We assume that the return is decided by

$$m(D_i,\varepsilon_i)-c(D_i,Z_i,\eta_i),$$

where m() is her expected income, c() refers to the cost of attending college, Z_i represents observable exogenous factors that affect the cost (proximity to hometown), ε_i and η_i are unobservable factors.

We are interested in the relationship between expected income and college education, Y_i = m(D_i, ε_i).

IV in structural models

D_i is an endogenous choice as

$$D_i = rg\max_d [m(d,arepsilon_i) - c(d,Z_i,\eta_i)].$$

• We can write $D_i = g(Z_i, \nu_i)$, where $\nu_i = c(\varepsilon_i, \eta_i)$.

Now, we have a "triangular system:"

 $Y_i = m(D_i, \varepsilon_i),$ $D_i = g(Z_i, \nu_i),$ $Z_i \perp \nu_i, \varepsilon_i \not\perp \nu_i.$

Ususally we assume g() is monotonic:

If $g(Z_i, \nu_i) > g(Z_i^{'}, \nu_i)$, then $g(Z_i, \nu_i^{'}) > g(Z_i^{'}, \nu_i^{'})$.

IV in structural models

- We call Z_i in the triangular system an instrumental variable.
- Like the supply shock in the fishing example, it provides exogenous variations for us to identify parameters in the system.
- The triangular system can incorporate many other scenarios and allows for arbitrary heterogeneity in effects.
- Both ε_i and ν_i might be high-dimensional.
- There is no restriction on whether the variables are continuous or discrete.
- This framework is built upon economic theories and captures complexities in reality.
- ► Non-compliance is not mentioned in such a system.

IV in linear models

- It is impossible to identify any causal parameter of interest in the triangular system without more structural restrictions.
- ▶ Let's consider the simplest scenario: both m() and g() are linear functions with homogeneous effects and no intercept, then

$$Y_{i} = \tau D_{i} + \varepsilon_{i},$$

$$D_{i} = \delta Z_{i} + \nu_{i},$$

$$Z_{i} \perp \nu_{i}, \varepsilon_{i} \not\perp \nu_{i}.$$

Cov(D_i, ε_i) ≠ 0, hence regressing Y_i on D_i leads to bias.
 But we can still estimate τ using a two-step approach.

- First, we estimate the second equation with OLS and obtain $\hat{\delta}$.
- This is known as the "first stage" regression.
- Then, note that

$$Y_{i} = \tau D_{i} + \varepsilon_{i}$$

= $\tau (\delta Z_{i} + \nu_{i}) + \varepsilon_{i}$
= $\xi Z_{i} + \tilde{\varepsilon}_{i}$,

where $\xi = \tau \delta$ and $Z_i \perp \tilde{\varepsilon}_i = \tau \nu_i + \varepsilon_i$.

- Hence, regressing Y_i on Z_i leads to a consistent estimate of $\tau \delta$.
- This is known as the "reduced-form" regression.
- The ratio of the two regression estimates is consistent for τ .
- ▶ This algorithm is known as the two-stage least squares (2SLS).

- There are two equivalent approaches to implement the 2SLS.
- The first stage is always necessary.
- Let's denote the predicted value of D_i as \hat{D}_i and the regression residual as $\hat{\nu}_i$.
- One approach is to regress Y_i on \hat{D}_i (the second stage).
- The OLS estimate will be consistent for τ .
- Consider the matrix form of the equations:

 $\begin{aligned} \mathbf{Y} &= \mathbf{D}\tau + \varepsilon, \\ \mathbf{D} &= \mathbf{Z}\delta + \nu. \end{aligned}$

Remember that

$$\begin{split} \hat{\delta} &= (\mathbf{Z}'\mathbf{Z})^{-1}(\mathbf{Z}'\mathbf{D}), \\ \widehat{\mathbf{D}} &= \mathbf{Z}(\mathbf{Z}'\mathbf{Z})^{-1}(\mathbf{Z}'\mathbf{D}) = \mathbf{P}_{\mathbf{Z}}\mathbf{D}, \\ \hat{\nu} &= \mathbf{D} - \widehat{\mathbf{D}} = (\mathbf{I} - \mathbf{P}_{\mathbf{Z}})\mathbf{D}. \end{split}$$

• The OLS estimate from regressing **Y** on $\widehat{\mathbf{D}}$ equals:

$$\begin{aligned} & \left(\widehat{\mathbf{D}}' \widehat{\mathbf{D}} \right)^{-1} \left(\widehat{\mathbf{D}}' \mathbf{Y} \right) \\ &= \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} \mathbf{P}_{\mathbf{Z}} \mathbf{D} \right)^{-1} \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} \mathbf{Y} \right) \\ &= \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} \mathbf{D} \right)^{-1} \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} (\mathbf{D} \tau + \varepsilon) \right) \\ &= \tau + \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} \mathbf{D} \right)^{-1} \left(\mathbf{D}' \mathbf{P}_{\mathbf{Z}} \varepsilon \right) \to \tau. \end{aligned}$$

It is natural to see that

$$Var[\hat{\tau}] = \left(\mathbf{D}'\mathbf{P}_{\mathbf{Z}}\mathbf{D}\right)^{-1} \left(\mathbf{D}'\mathbf{P}_{\mathbf{Z}}\varepsilon\varepsilon'\mathbf{P}_{\mathbf{Z}}\mathbf{D}\right) \left(\mathbf{D}'\mathbf{P}_{\mathbf{Z}}\mathbf{D}\right)^{-1}$$

- It is more complicated than the variance for the OLS estimator but also takes the sandwich form.
- Intuitively, we first project D_i onto the space spanned by Z_i .
- Next, we project Y_i onto the same space, with the projected D_i as the bases.
- We can estimate the variance by replacing ε with the regression residuals in the second stage.

Control function

- The second approach is known as the "control function" approach in econometrics.
- In the second stage, we regress Y_i on D_i and $\hat{\nu}_i$.
- We can show that the estimated coefficient for D_i will be consistent for τ.
- Note that ν_i is the endogenous part in D_i and ν̂_i is unbiased for ν_i.
- If we can control for the endogenous part, the estimation will be unbiased.
- The three approaches give you numerically equivalent results when the models are correct.

Control function: application

- ## The OLS estimate is 3.886754
- ## The 2SLS estimate is 2.964144
- ## The 2SLS estimate is 2.964144
- ## The control function estimate is 2.964144

2SLS with covariates

- When there are covariates, we treat each covariate as its own instrument.
- The two regression models are

$$\begin{split} \mathbf{Y} &= \tilde{\mathbf{X}} \boldsymbol{\beta} + \boldsymbol{\varepsilon}, \\ \tilde{\mathbf{X}} &= \tilde{\mathbf{Z}} \boldsymbol{\gamma} + \boldsymbol{\nu}, \end{split}$$

where
$$\tilde{\mathbf{X}} = (\mathbf{D}, \mathbf{X})$$
 and $\tilde{\mathbf{Z}} = (\mathbf{Z}, \mathbf{X})$.

• We know that $\tilde{\mathbf{X}} = \mathbf{P}_{\tilde{\mathbf{Z}}}\tilde{\mathbf{X}}$.

► Hence,

$$\begin{split} \hat{\boldsymbol{\beta}} &= \left(\hat{\tilde{\mathbf{X}}}'\hat{\tilde{\mathbf{X}}}\right)^{-1} (\hat{\tilde{\mathbf{X}}}'\mathbf{Y}) \\ &= \left(\tilde{\mathbf{X}}'\mathbf{P}_{\tilde{\mathbf{Z}}}\tilde{\mathbf{X}}\right)^{-1} \left(\tilde{\mathbf{X}}'\mathbf{P}_{\tilde{\mathbf{Z}}}'\mathbf{Y}\right). \end{split}$$

Generalized methods of moments

- Sometimes we have multiple instruments for the treatment.
- Each instrument provides an independent source of exogenous variation.
- This is a scenario known as "over-identification."
- We can combine these instruments together via the generalized methods of moments (GMM).
- Each instrument variable provides a moment condition, $\Psi_k(\beta)$.
- We try to find $\hat{\beta}$ such that

$$\hat{eta} = rgmin_{eta} \hat{E}[\Psi]' \widehat{Var}^{-1}[\Psi] \hat{E}[\Psi].$$

 The theory on GMM is well documented by Newey and McFadden (1994).

- We have reviewed the literature on instrumental variables following the econometric tradition.
- We didn't not use the potential outcome notations at all!
- Are the 2SLS estimates causal from the design-based perspective?
- Note that the triangular system satisfies Assumptions 1-4 for identifying the LATE.
- Random assignment results from that $Z_i \perp \nu_i$.
- Exclusion restriction and the first stage hold due to the functional form.
- Monotonicity is assumed for g().

- ► Angrist, Imbens, and Rubin (1996) first showed that when 1) D_i and Z_i are binary, and 2) the effects are heterogeneous, the 2SLS estimator \$\tau_{2SLS}\$ equals the Wald estimator \$\tau_{Wald}\$.
- Compliers are the agents who are encouraged to select into the treatment by the instrument.
- We start from an economic model and end up with an interpretation rooted in the potential outcome framework.
- It suggests that the economic approach and the statistical approach are capturing the same concepts.
- We can justify the potential outcome framework with social science theory, or analyze social science problems with the idea of counterfactual.
- This finding won Angrist and Imbens a Nobel prize.

In this case, the triangular system boils down to:

$$Y_i = \tau_i D_i + \varepsilon_i,$$

$$D_i = \mathbf{1}\{(\delta_i Z_i + \nu_i) \ge 0\}.$$

- We can see that $D_i(0) = \mathbf{1}\{\nu_i \ge 0\}$ and $D_i(1) = \mathbf{1}\{\delta_i + \nu_i \ge 0\}.$
- δ_i and ν_i determine the response of unit *i* to Z_i hence *i*'s type.
- For example, compliers are those with $-\delta_i \leq \nu_i < 0$.
- The analysis in Angrist, Imbens, and Rubin (1996) indicates that the 2SLS estimator's result is completely driven by these units.

2SLS: application

The LATE is 2.719

- ## Estimate from the Wald estimator is 2.455
- ## The 2SLS estimate is 2.455
- ## SE of the 2SLS estimate is 0.559

- These results led to the credibility revolution in the 90s.
- Using IVs in observational studies to identify the LATE became a fad.
- Angrist (1990): draft lottery in the Vietnam war veteran status - earnings on the labor market.
- Angrist and Krueger (1991): birth season age when dropping out of high school - earnings on the labor market.
- Acemoglu, Johnson, and Robinson (2001): settlers' mortality inclusive institutions - economic development.
- But many applications turned out to be not as credible as we thought.

References I

- Acemoglu, Daron, Simon Johnson, and James A Robinson. 2001.
 "The Colonial Origins of Comparative Development: An Empirical Investigation." *American Economic Review* 91 (5): 1369–1401.
- Angrist, Joshua D. 1990. "Lifetime Earnings and the Vietnam Era Draft Lottery: Evidence from Social Security Administrative Records." *The American Economic Review*, 313–36.
- Angrist, Joshua D, Guido W Imbens, and Donald B Rubin. 1996. "Identification of Causal Effects Using Instrumental Variables." *Journal of the American Statistical Association* 91 (434): 444–55.
- Angrist, Joshua D, and Alan B Krueger. 1991. "Does Compulsory School Attendance Affect Schooling and Earnings?" *The Quarterly Journal of Economics* 106 (4): 979–1014.
 Newey, Whitney K, and Daniel McFadden. 1994. "Large Sample Estimation and Hypothesis Testing." *Handbook of Econometrics* 4: 2111–2245.