Regression Discontinuity Design I

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Linear Methods in Causal Inference POL1784

Review

- Assumptions underlying the instrumental variable method are hard to satisfy in observational studies.
- If an instrument is weak, then the finite-sample bias of the estimate will be large and the confidence interval will not have the correct coverage rate.
- We can use the Anderson-Rubin test or adjust the critical value of the t-statistic based on the value of the first-stage F-statistic.
- ▶ We can test exclusion restriction and monotonicity jointly.
- Under model-specification assumptions, we can charaterize the compliers and generalize the LATE to the ATE.

Some history

- Regression discontinuity design (RDD) became popular in political science since Lee (2008) applied this method to analyze congressional elections in the US.
- Thistlethwaite and Campbell (1960) first developed this idea in psychology.
- ► Hahn, Todd, and Van der Klaauw (2001) formally discussed the identification assumptions in RDD.
- Porter (2003) proposed the kernel regression estimator based on the work by Fan and Gijbels (1996).
- Imbens and Kalyanaraman (2012) introduced the first data-driven bandwidth selector.
- Calonico, Cattaneo, and Titiunik (2014) derived the asymptotic distributions for a group of RDD estimators.

Setup

- We possess a dataset with N units.
- We observe (Y_i, D_i, Z_i) for each unit $i \in \{1, 2, \dots, N\}$.
- ► Z_i is called the running variable, which decides the value of the binary treatment D_i.
- We denote Z_i 's density function as f(z).
- ► There exists a cutoff c (often set to be 0) such that D_i = 1{Z_i ≥ c}.
- This setup is known as the sharp RDD.
- The causal parameter of interest, τ_{SRD} , is defined as

$$E[Y_i(1) - Y_i(0)|Z_i = c].$$

- Notice that the causal parameter is conditional/local by definition.
- It only captures the effect on those with $Z_i = c$.

Ideal experiment behind sharp RDD

- This setup is slightly different from what we saw before.
- Positivity is violated everywhere but at the cutoff point.
- ▶ We can treat RDD as a randomized experiment conducted in the block defined by Z_i = c.
- Suppose there are 1,000 congressional elections, in all of which both candidates win 50% of the votes.
- ▶ We then randomly select a winner for each election with a coin flip, which implies that $D_i \perp \{Y_i(1), Y_i(0)\}|Z_i = c$.
- In theory, all we need to do is to calculate the difference in means.
- The result reflects the causal effect of the winner's attributes (e.g., party affiliation).

Ideal experiment behind sharp RDD

- Unfortunately, we do not have that many elections with a 50-50 split in practice.
- We need structural restrictions on the outcome such that we can approximate the two average outcomes using data beyond the cutoff point.
- The RDD estimate is thus biased by default.
- Yet the bias diminishes to zero as the sample size increases, as we can use more observations that are close to the cutoff.
- Then, the estimate converges to the unbiased one under the ideal experiment.

Identification in sharp RDD

- Identification in the setting of sharp RDD requires the continuity of the expected outcome (Hahn, Todd, and Van der Klaauw 2001).
- Define $\mu(z) = E[Y_i | Z_i = z]$, $\mu_+ = \lim_{z \to c^+} \mu(z)$, and $\mu_- = \lim_{z \to c^-} \mu(z)$.
- Continuity means that $\mu_+ = E[Y_i(1)|Z_i = c]$ and $\mu_- = E[Y_i(0)|Z_i = c]$.
- If so, we have

$$\tau_{SRD} = \mu_+ - \mu_-.$$

• All we need to do is to estimate μ_+ and μ_- .

Estimation in sharp RDD

- ► To estimate µ₊ and µ₋, the most common choice is the kernel regression estimator.
- ► Given a selected bandwidth h, we can fit the model within the windows [c h, c] and [c, c + h], respectively:

$$(\hat{\mu}_+, \hat{eta}_+)$$

= arg min $\sum_{\mu, eta}^N \mathbf{1}\{Z_i \ge c\} (Y_i - \mu - eta(Z_i - c))^2 \, \mathsf{K}\left(rac{Z_i - c}{h}
ight)$

$$\hat{\tau}_{SRD} = \hat{\mu}_+ - \hat{\mu}_-.$$

Estimation in sharp RDD

- Remember that to improve the precision of prediction, units that are closer to the cutoff should receive a larger weight.
- We usually use the triangular kernel in practice.



Sharp RDD: application

- ▶ We will use Meyersson (2014) as an illustrative example.
- The author investigated the consequences of the 1994 municipal elections in Turkey.
- Does the victory of Islamic candidates endanger the rights of women?
- In the raw data, we can find a negative correlation between the vote share of Islamic candidates and high school attainment of women across Turkish cities.
- But results from RDD tell a different story.
- The victory of Islamic mayors actually empowered "the poor and pious" and encouraged women to receive more education.

Sharp RDD: application



Sharp RDD: application

Sharp RD estimates using local polynomial regression. ## ## Number of Obs. 2630 ## BW type mserd ## Kernel Triangular ## VCE method NN ## ## Number of Obs. 2315 315 ## Eff. Number of Obs. 529 266 ## Order est. (p) 1 1 ## Order bias (q) 2 2 ## BW est. (h) 0.172 0.172## BW bias (b) 0.286 0.286 ## rho (h/b) 0.603 0.603 ## Unique Obs. 2313 315 ## ## _____ ## Method Coef. Std. Err. $P > |_{z}|_{5}$ z

Implement RDD in practice

- It is always critical to draw plots in RDD.
- The result would not be convincing if we cannot see the jump of the outcome variable from the plot.
- Note that the plot is based on global polynomials rather than kernel regressions.
- It does not include all the data points.
- Never use global polynomials to estimate the two intercepts (Gelman and Imbens 2019).
- We are interested in local quantities rather than global fitness.
- Estimates of the intercepts may be driven by points that are far away from the cutoff if you use global polynomials.

Implement RDD in practice

- In practice, some scholars select the bandwidth using rdrobust and then manually fit two regressions within the selected window.
- What is the problem of this approach?
- It is inconsistent with the bandwidth selector and often leads to larger biases.
- Bias correction is ignored and inference may be problematic.
- ► One may add higher order terms of (Z_i − c) into the regressions.
- But for sharp RDD, linear regression has the optimal rate of convergence due to its nice performance on the boundary (Porter 2003).

Test the assumption of continuity

- As continuity is crucial for identification in sharp RDD, we should test its validity in practice.
- A common approach is to rely on placebo tests.
- Several ways to do this:
 - Apply the estimator to a covariate;
 - Apply the estimator to a placebo outcome;
 - Apply the estimator at a point other than the cutoff (placebo treatment).
- Meyersson (2014) did a really good job.

Tests in sharp RDD: application



Tests in sharp RDD: application

##		CI Lower	CI Upper
##	Conventional	-0.01464976	0.04035605
##	Bias-Corrected	-0.01486825	0.04013756
##	Robust	-0.02051897	0.04578828
##		CI Lower	CI Upper
##	Conventional	-0.02616607	0.03410509
##	Bias-Corrected	-0.02929148	0.03097967
##	Robust	-0.03526899	0.03695719
##		CI Lower	CI Upper
## ##	Conventional	CI Lower -0.02713105	CI Upper 0.02631536
## ## ##	Conventional Bias-Corrected	CI Lower -0.02713105 -0.02997479	CI Upper 0.02631536 0.02347162
## ## ## ##	Conventional Bias-Corrected Robust	CI Lower -0.02713105 -0.02997479 -0.03497856	CI Upper 0.02631536 0.02347162 0.02847539
## ## ## ##	Conventional Bias-Corrected Robust	CI Lower -0.02713105 -0.02997479 -0.03497856 CI Lower	CI Upper 0.02631536 0.02347162 0.02847539 CI Upper
## ## ## ## ##	Conventional Bias-Corrected Robust Conventional	CI Lower -0.02713105 -0.02997479 -0.03497856 CI Lower -0.01436241	CI Upper 0.02631536 0.02347162 0.02847539 CI Upper 0.03079918
## ## ## ## ##	Conventional Bias-Corrected Robust Conventional Bias-Corrected	CI Lower -0.02713105 -0.02997479 -0.03497856 CI Lower -0.01436241 -0.01068460	CI Upper 0.02631536 0.02347162 0.02847539 CI Upper 0.03079918 0.03447699

Test the assumption of continuity

- The assumption is also violated when units in the sample self-select into one side of the cutoff (known as sorting).
- For example, students may cheat to meet the requirement of a scholarship.
- ► As a result, the density function f(z) will not change smoothly across the cutoff.



Test the assumption of continuity

- McCrary (2008) develop the first formal test of the discontinuity in density.
- Essentially, we are testing whether

$$heta = \ln \lim_{z o 0^+} f(z) - \ln \lim_{z o 0^-} f(z)$$

deviates significantly from 0.

- Similarly, we estimate the two boundary points of ln f(z) using local regression (Y_i is not needed).
- Cattaneo, Jansson, and Ma (2020) provide an augmented algorithm for non-parametric density estimation.

Tests in sharp RDD: application



- Let's introduce some notations for deriving the bias of the kernel regression estimator.
- Denote $(Y_1, Y_2, ..., Y_N)'$ as **Y**, $(Z_1, Z_2, ..., Z_N)'$ as **Z**, and

$$\mathbf{R} = (\iota, \mathbf{Z})$$
$$\mathbf{W}_{+} = \left\{ \mathbf{1} \{ Z_{i} \ge c \} K \left(\frac{Z_{i} - c}{h_{N}} \right) \right\}_{N \times N}$$
$$\mathbf{M} = (\mu(Z_{1}), \mu(Z_{2}), \dots, \mu(Z_{N}))'$$

where ι is a vector with N 1s and W is a diagonal matrix of weights.

• We use $\mu_{+}^{(k)}$ to denote the *k*-th order derivative of μ_{+} (similar for μ_{-}) and $\sigma^{2}(z)$ to denote $Var[Y_{i}|Z_{i} = z]$.

- Notice that $\mathbf{Y} = \mathbf{M} + \varepsilon$ with $E[\varepsilon_i | Z_i] = 0$.
- The estimate $\hat{\mu}_+$ equals to the first row of

$$(\mathsf{R}'\mathsf{W}_{+}\mathsf{R})^{-1}(\mathsf{R}'\mathsf{W}_{+}\mathsf{Y})$$

=(\epsilon'\epsilon_{+}\epsilon)^{-1}(\epsilon'\epsilon_{+}\epsilon) + (\epsilon'\epsilon_{+}\epsilon)^{-1}(\epsilon'\epsilon_{+}\epsilon) + (\epsilon'\epsilon_{+}\epsilon_{+}\epsilon)^{-1}(\epsilon'\epsilon_{+}\epsilon) + (\epsilon'\epsilon_{+}\e

► Expectation of the second term is zero and we have the Taylor expansion for µ(Z_i):

$$\mu(Z_i) = \mu_+(0) + \mu_+^{(1)}(0)Z_i + \frac{\mu_+^{(2)}(0)}{2}Z_i^2 + \nu_i$$

Hence,

$$\mathbf{M} = \mathbf{R} \begin{pmatrix} \mu_+(0) \\ \mu_+^{(1)}(0) \end{pmatrix} + \mathbf{S}_2 \frac{\mu_+^{(2)}(0)}{2} + \nu$$

where $\mathbf{S}_2 = (Z_1^2, Z_2^2, \dots, Z_N^2)$ and $\nu = (\nu_1, \nu_2, \dots, \nu_N)$.

▶ Now, the estimation bias of $\hat{\mu}_+$, $E[\hat{\mu}_+] - \mu_+$, is the first row of

$$(\mathbf{R}'\mathbf{W}_{+}\mathbf{R})^{-1}\left(\mathbf{R}'\mathbf{W}_{+}\mathbf{S}_{2}\frac{\mu_{+}^{(2)}(0)}{2}\right) + (\mathbf{R}'\mathbf{W}_{+}\mathbf{R})^{-1}\left(\mathbf{R}'\mathbf{W}_{+}\nu\right)$$

- The convergence rates of these two terms rely on the properties of the kernel.
- Via some cumbersome calculation, we can see that

$$E[\hat{\mu}_+] - \mu_+ = C_1 \mu_+^{(2)}(0)h^2 + o_p(h^2)$$

▶ We can similarly derive the variance of $\hat{\mu}_+$ using the properties of regression:

$$Var[\hat{\mu}_+] = \frac{C_2}{Nh} \frac{\sigma_+^2(0)}{f_+(0)} + o_p\left(\frac{1}{Nh}\right)$$

- Obviously, the bias and the variance of $\hat{\mu}_{-}$ have similar forms.
- ► More generally, we can estimate the k-th order derivative of µ₊ and µ₋ with a p-th order local regression.
- The bias will have an order of p + 1.

Bandwidth selection in RDD

- Before Imbens and Kalyanaraman (2012), the practice is to minimize the regression's MSE on the entire support of Z using cross-validation.
- Imbens and Kalyanaraman (2012) argue that we should select a bandwidth to minimize the MSE of the estimator:

$$\begin{split} \mathsf{MSE}(h) = & \mathbb{E} \left[\hat{\tau}_{SRD} - \tau_{SRD} | \mathbf{Z} \right]^2 \\ = & (\mathbb{E} \left[\hat{\tau}_{SRD} | \mathbf{Z} \right] - \tau_{SRD})^2 + \mathsf{Var} \left[\hat{\tau}_{SRD} | \mathbf{Z} \right] \\ = & \mathsf{Bias}^2 + \mathsf{Variance.} \end{split}$$

Bandwidth selection for optimizing MSE

Imbens and Kalyanaraman (2012) show that in practice we can minimize the asymptotic MSE:

$$AMSE(h) = C_1 h^4 \left(\mu_+^{(2)}(0) - \mu_-^{(2)}(0) \right)^2 + \frac{C_2}{Nh} \frac{\sigma^2(0)}{f(0)}$$

From the expression we can solve the optimal bandwidth:

$$h_{N}^{*} = C \left(\frac{\frac{\sigma^{2}(0)}{f(0)}}{\left(\mu_{+}^{(2)}(0) - \mu_{-}^{(2)}(0) \right)^{2}} \right)^{\frac{1}{5}} N^{-\frac{1}{5}}.$$

▶ In practice, we can estimate h_N^* with a plug-in estimator.

- Imbens and Kalyanaraman (2012) do not prove the asymptotic normality of the RDD estimate.
- Calonico, Cattaneo, and Titiunik (2014) study the asymptotic distribution of the studentized RDD estimate:

$$\frac{\hat{\tau}_{SRD} - \tau_{SRD}}{\sqrt{Var[\hat{\tau}_{SRD}]}}.$$

- They show that the bandwidth selected via the previous algorithm is too wide to guarantee the the asymptotic normality of the estimate.
- We need $h = o_p \left(N^{-\frac{1}{5}} \right)$ while the algorithm leads to $h_N^* = O_p \left(N^{-\frac{1}{5}} \right)$.
- Consequently, the studentized estimate will be asymptotically biased.

Intuitively,

$$\frac{\hat{\tau}_{SRD} - \tau_{SRD}}{\sqrt{Var[\hat{\tau}_{SRD}]}} = \frac{\hat{\tau}_{SRD} - E[\hat{\tau}_{SRD}]}{\sqrt{Var[\hat{\tau}_{SRD}]}} + \frac{E[\hat{\tau}_{SRD}] - \tau_{SRD}}{\sqrt{Var[\hat{\tau}_{SRD}]}}$$

- ► The first term is a weighted average of residuals and converges to N(0,1) by CLT.
- We need to guarantee that the second term is sufficiently small.
- Remember that the numerator is $O_p(h^2)$ and the denominator is $O_p(\frac{1}{\sqrt{Nh}})$, thus the total bias is $O_p(\sqrt{Nh^5})$ and does not decline to zero.

- Calonico, Cattaneo, and Titiunik (2014) provide a bias-correction estimator
 ^{bc}_{SRD}.
- Intuitively, we use another local regression estimator with bandwidth b_N to estimate the second-order derivative of μ_+ and μ_- and subtract them from $\hat{\mu}_+$ and $\hat{\mu}_-$.
- Bias correction introduces extra uncertainty (from the extra local regression) into the estimate, hence the variance has to be adjusted accordingly.
- The bias-corrected CI does not account for this extra uncertainty.
- They propose two variance estimators, one based on regression analysis and the other based on the idea of nearest neighborhood matching (Abadie and Imbens 2006).

Calonico, Cattaneo, and Titiunik (2014) prove that

$$\frac{\hat{\tau}_{SRD}^{bc} - \tau_{SRD}}{\sqrt{Var[\hat{\tau}_{SRD}^{bc}]}} \rightarrow N(0, 1)$$

as long as $N\min\{b^5, h^5\}\max\{b^2, h^2\} \to 0.$

- In other words, we can still use the algorithm in Imbens and Kalyanaraman (2012) to select the bandwidth.
- We just need to modify the obtained estimate to ensure asymptotic normality.

Bandwidth selection for the bias-correction estimator

- ▶ Notice that now we need two bandwidths, *h* and *b*.
- In practice, we start from two pilot bandwidths h₀ and b₀ selected by some rule of thumb.
- ▶ We use h₀ and b₀ to run the bias-correction estimator for the second-order derivatives and obtain b^{*} using the algorithm in Imbens and Kalyanaraman (2012).
- ► We then use h₀ and b^{*} to run the bias-correction estimator for the two intercepts and obtain h.

Covariates in RDD

- In theory, there should be no confounder in RDD as it approximates a simple experiment.
- But we can use the information contained in covariates to enhance the estimator's efficiency.
- Calonico et al. (2019) discuss how to include covariates in the estimation.
- The basic idea is just the FWL theorem for regression.

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