Panel Data Analysis III

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Linear Methods in Causal Inference POLI784

Review

- ▶ The previous class started from introducing the DID estimator.
- It is motivated by the TWFE model when the data have a DID structure.
- ▶ But the estimator only needs the assumption of parallel trends and is robust to heterogeneous treatment effects.
- ▶ This is not true when the data have a more complex structure.
- The within estimator may generate negative weights for individualistic treatment effects.
- ▶ We can fix the problem by not using untreated observations when fitting the model.
- ▶ This leads to the idea of counterfactual estimation.

The classic factor model

- ▶ Remember that in the TWFE model, we assume that $h_t(\mathbf{U}_i) = \mu + \alpha_i + \xi_t$.
- ▶ Bai (2003) relaxes this assumption to $h_t(\mathbf{U}_i) = \mathbf{f}_t' \lambda_i$.
- ▶ Both \mathbf{f}_t and λ_i are r-dimensional vectors.
- ▶ The number of parameters is $(N + T) * r \le N * T$.
- The former is known as factors, and the latter as factor loadings.
- ▶ When $\mathbf{f}_t = (1, \xi_t)'$ and $\lambda_i = (\alpha_i, 1)'$, it boils down to the TWFE model.
- Such a model captures the interaction between variables that vary only over time and only across units.
- ▶ E.g., country *i*'s endowments in *r* different resources and the prices of them in year *t*.

The classic factor model

▶ Let's omit the covariates, then

$$Y_{it} = \mathbf{f}_{t}' \lambda_{i} + \varepsilon_{it},$$

▶ We can write the model in matrices:

$$\mathbf{Y} = \mathbf{F} \mathbf{\Lambda}' + \varepsilon.$$

where **Y** is a $T \times N$ matrix; **F** $(T \times r)$ represents factors and Λ $(N \times r)$ represents factor loadings.

- ▶ Bai (2003) shows that **F** can be found via eigenvalue decomposition of the matrix **YY**'/T.
- We find a matrix F such that

$$\left(\frac{1}{NT}\sum_{i=1}^{N}\mathbf{Y}_{i}\mathbf{Y}_{i}^{\prime}\right)\hat{\mathbf{F}}=\hat{\mathbf{F}}\mathbf{V}_{NT},$$

where \mathbf{V}_{NT} is the matrix of eigenvalues.

► Then, we can estimate Λ via OLS, treating each $\hat{\mathbf{f}}_t$ as a known value.

The classic factor model

▶ Bai (2009) considers the DGP with covariates:

$$\mathbf{Y} = \mathbf{X}\beta + \mathbf{F}\Lambda' + \varepsilon.$$

▶ Knowing β , we can calculate residuals and estimate **F** and Λ as before; knowing **F** and Λ , we can estimate β via

$$\hat{\beta} = \left(\sum_{i=1}^{N} \mathbf{X}_{i}' \mathbf{Q}_{\hat{\mathbf{F}}} \mathbf{X}_{i}\right)^{-1} \left(\sum_{i=1}^{N} \mathbf{X}_{i}' \mathbf{Q}_{\hat{\mathbf{F}}} \mathbf{Y}_{i}\right)$$

where $\mathbf{Q}_{\hat{\mathbf{r}}}$ is the residual-making matrix of $\hat{\mathbf{F}}$.

- ▶ We can start from some initial values and iterate this process until convergence.
- ▶ Bai shows that $\sqrt{NT}(\hat{\beta} \beta)$ converges to a normal distribution under regularity conditions.

- Abadie, Diamond, and Hainmueller (2010) propose the synthetic control (SC) method for case studies.
- ▶ It assumes that there are N untreated units and a single treated unit (i = 1).
- ▶ There are T_0 pre-treatment periods and T_1 post-treatment periods.
- The untreated potential outcome obeys the classic factor model as in Bai (2003) or Bai (2009).
- ▶ We are interested in is the treatment effect on unit 1 in any period $t > T_0$:

$$\tau_{1t} = Y_{1t}(1) - Y_{1t}(0).$$

- Abadie, Diamond, and Hainmueller (2010) show that we don't need to estimate factors or factor loadings directly in this case.
- ▶ We instead attempt to find a group of weights $\{w_i\}_{i=2}^{N+1}$ such that

$$\sum_{i=2}^{N+1} w_i = 1, \sum_{i=2}^{N+1} w_i Y_{it} = Y_{1t}$$

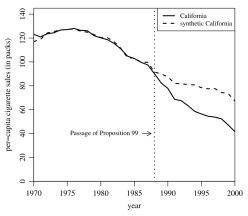
for any $1 \le t \le T_0$.

- We require the treated unit to lie in the "convex hull" of the untreated ones.
- ▶ The predicted counterfactual for unit 1 in period $t > T_0$ is

$$\hat{Y}_{1t}(0) = \sum_{i=2}^{N+1} w_i Y_{it}.$$

▶ Hence, we can estimate τ_{1t} by $\hat{\tau}_{1t} = Y_{1t}(1) - \hat{Y}_{1t}(0)$.

▶ Ideally, we would like to see a picture like this:



- ▶ Balancing pre-treatment outcomes can effectively remove interactive fixed effects hence $\hat{\tau}_{1t}$ is consistent for τ_{1t} .
- ▶ But $\hat{\tau}_{1t}$ has no asymptotic distribution with only one treated observation.
- ▶ We rely on a permutation test for statistical inference.
- ▶ It is similar to a placebo test: We replace the treated unit with a randomly selected unit from the control group and estimate the effect.
- This is not Fisher's randomization test since we do not know the treatment assignment algorithm.
- It assumes that each unit has the same probability of being treated.
- ► There are approaches to construct non-asymptotic confidence intervals but they tend to be wide.

Generalized synthetic control

- ▶ In practice, most data include more than one treated unit.
- ▶ In theory, we can find a unique set of weights for each treated unit.
- But it could be more straightforward to estimate factors and factor loadings.
- ▶ Xu (2017) considers the following model:

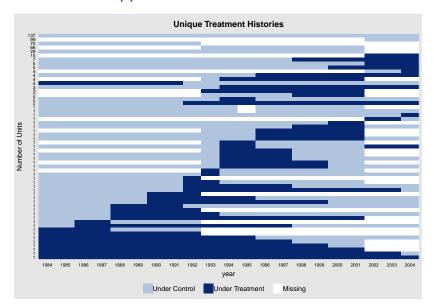
$$Y_{it}(0) = \mathbf{f}'_t \lambda_i + \mathbf{X}'_{it} \beta + \varepsilon_{it},$$

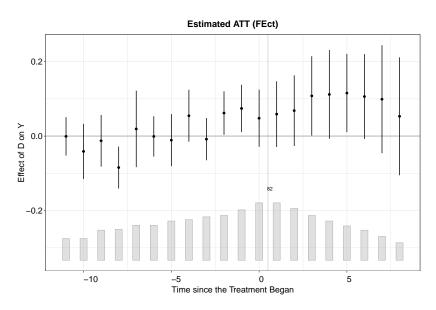
$$Y_{it}(1) = Y_{it}(0) + \tau_{it}.$$

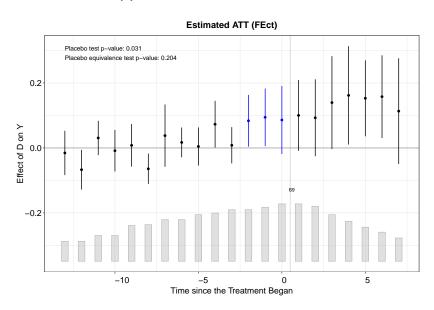
- As in counterfactual estimation, we estimate \mathbf{f}_t , λ_i , and β using untreated observations.
- Then, we generate predictions of the counterfactual and estimate the ATT with:

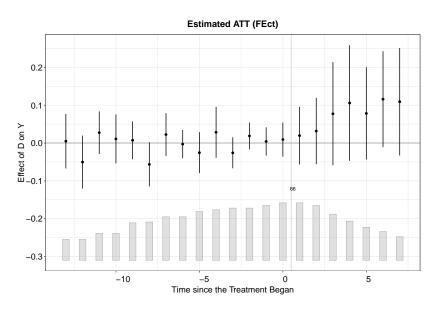
$$\widehat{ au}_{ATT} = rac{1}{|\mathcal{M}|} \sum_{(i,t) \in \mathcal{M}} \left(Y_{it} - \hat{Y}_{it}(0)
ight).$$

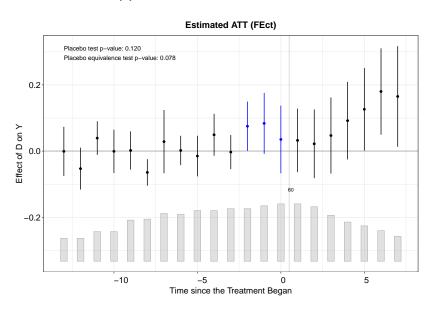
- ▶ Let's examine the study in Anzia and Berry (2011).
- ► The argument is that female legislators face more structural barriers than their male counterparts.
- Once elected, they will have better performance.
- ▶ The data include 700 districts in the US over 21 years.
- ▶ $D_{it} = 1$ if there is a female legislator from district i in year t.
- ▶ The outcome is measured by the amount of federal spending a legislator can secure for their district in year *t*.

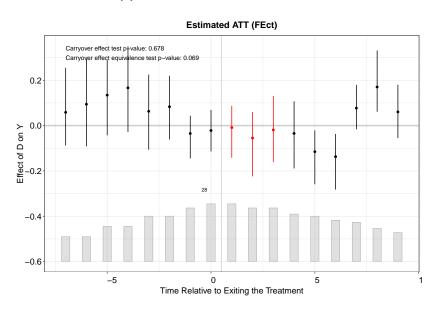












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