

Instrumental Variable II

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Linear Methods in Causal Inference
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Review

- ▶ We discussed how to identify causal effects when there exists non-compliance.
- ▶ In this case, treatment assignment Z_i no longer equals treatment status D_i .
- ▶ We can always identify the intention-to-treat effect but may care about the local average treatment effect (LATE), or the effect on the compliers.
- ▶ It can be identified if 1. Z_i is randomly assigned; 2. it affects Y_i only through D_i ; 3. it changes the value of D_i monotonically.
- ▶ Then, we can estimate the LATE using the Wald estimator.

History of IV

- ▶ We called treatment assignment Z_i under non-compliance an instrumental variable (IV).
- ▶ But the idea of using an IV for causal inference was proposed in a very different context.
- ▶ It was introduced by economists to study simultaneous structural equations.
- ▶ People used to believe that we can describe a social system (e.g., the US economy) with a large number of equations.
- ▶ Variables on the left-hand side are referred to as endogenous, while those only appear on the right-hand side are called exogenous.
- ▶ Identification means that we can estimate coefficients in these models consistently.
- ▶ Instrumental variables were proposed as a solution to the identification problem.

History of IV

- ▶ An economist observes the everyday price and quantity of transaction for fish in a market over N days: $(P_i, Q_i)_{i=1}^N$.
- ▶ She wants to identify the demand curve of fish: $P_i = d(Q_i)$.
- ▶ But we only have the price and quantity at equilibrium, which is also affected by the supply curve: $P_i = s(Q_i)$.
- ▶ Suppose both the demand and the supply curves are linear:

$$\begin{aligned}P_{di} &= a_d - b_d Q_{di} + \varepsilon_{di}, \\P_{si} &= a_s + b_s Q_{si} + \varepsilon_{si}.\end{aligned}$$

- ▶ We know that at equilibrium,

$$\begin{aligned}P_i &= \frac{a_s b_d + a_d b_s + b_s(\varepsilon_{di} - \varepsilon_{si})}{b_d + b_s}, \\Q_i &= \frac{a_d - a_s + \varepsilon_{di} - \varepsilon_{si}}{b_d + b_s}.\end{aligned}$$

History of IV

- ▶ Even with a large sample, we can only estimate

$$\frac{a_s b_d + a_d b_s}{b_d + b_s} \text{ and } \frac{a_d - a_s}{b_d + b_s}$$

but not (a_d, b_d) .

- ▶ Suppose there is a shock Z_i on the supply side such as storms at sea:

$$P_{di} = a_d - b_d Q_{di} + \varepsilon_{di},$$

$$P_{si} = a_s + b_s Q_{si} + c_s Z_i + \varepsilon_{si},$$

$$Z_i \perp (\varepsilon_{di}, \varepsilon_{si}).$$

- ▶ Z_i is known as an instrumental variable.

History of IV

- ▶ Now, at equilibrium, we have

$$P_i = \frac{a_s b_d + a_d b_s + c_s b_d Z_i + b_s (\varepsilon_{di} - \varepsilon_{si})}{b_d + b_s},$$
$$Q_i = \frac{a_d - a_s - c_s Z_i + \varepsilon_{di} - \varepsilon_{si}}{b_d + b_s}.$$

- ▶ From these relationships, we can obtain estimates of the two slopes

$$\frac{c_s b_d}{b_d + b_s} \text{ and } \frac{-c_s}{b_d + b_s}.$$

- ▶ Their ratio allows us to identify the parameter b_d .

IV in structural models

- ▶ Let's consider an economic model of income and education.
- ▶ An individual i maximizes their return by deciding whether to attend college, $D_i \in \{0, 1\}$.
- ▶ We assume that the return is decided by

$$m(D_i, \varepsilon_i) - c(D_i, Z_i, \eta_i),$$

where $m()$ is their expected income, $c()$ refers to the cost of attending college, Z_i represents observable exogenous factors that affect the cost (proximity to hometown), ε_i and η_i are unobservable factors.

- ▶ We are interested in the relationship between expected income and college education, $Y_i = m(D_i, \varepsilon_i)$.

IV in structural models

- ▶ D_i is an endogenous choice as

$$D_i = \arg \max_d [m(d, \varepsilon_i) - c(d, Z_i, \eta_i)].$$

- ▶ We can write $D_i = g(Z_i, \nu_i)$, where $\nu_i = (\varepsilon_i, \eta_i)'$.
- ▶ Now, we have a “triangular system:”

$$Y_i = m(D_i, \varepsilon_i),$$

$$D_i = g(Z_i, \nu_i),$$

$$Z_i \perp \nu_i, \varepsilon_i \not\perp \nu_i.$$

- ▶ Usually we assume $g()$ is monotonic:

$$\text{If } g(Z_i, \nu_i) > g(Z'_i, \nu_i), \text{ then } g(Z_i, \nu'_i) > g(Z'_i, \nu'_i).$$

IV in structural models

- ▶ We call Z_i in the triangular system an instrumental variable.
- ▶ Like the supply shock in the fishing example, it provides exogenous variations for causal identification in the system.
- ▶ The triangular system can incorporate many other scenarios and allows for arbitrary heterogeneity in effects.
- ▶ Both ε_i and ν_i might be high-dimensional.
- ▶ There is no restriction on whether the variables are continuous or discrete.
- ▶ This framework is built upon economic theories and captures complexities in reality.
- ▶ Non-compliance is not mentioned in such a system.

IV in linear models

- ▶ It is impossible to identify any causal parameter of interest in the triangular system without more structural restrictions.
- ▶ Let's consider the simplest scenario: both $m()$ and $g()$ are linear functions with homogeneous effects and no intercept, then

$$Y_i = \tau D_i + \varepsilon_i,$$

$$D_i = \delta Z_i + \nu_i,$$

$$Z_i \perp \nu_i, \varepsilon_i \not\perp \nu_i.$$

- ▶ $\text{Cov}(D_i, \varepsilon_i) \neq 0$, hence regressing Y_i on D_i leads to bias.
- ▶ But we can still estimate τ using a two-step approach.

Two-stage least squares

- ▶ First, we estimate the second equation with OLS and obtain $\hat{\delta}$.
- ▶ This is known as the “first stage” regression.
- ▶ Then, note that

$$\begin{aligned}Y_i &= \tau D_i + \varepsilon_i \\&= \tau(\delta Z_i + \nu_i) + \varepsilon_i \\&= \xi Z_i + \tilde{\varepsilon}_i,\end{aligned}$$

where $\xi = \tau\delta$ and $Z_i \perp \tilde{\varepsilon}_i = \tau\nu_i + \varepsilon_i$.

- ▶ Hence, regressing Y_i on Z_i leads to a consistent estimate of $\tau\delta$.
- ▶ This is known as the “reduced-form” regression.
- ▶ The ratio of the two regression estimates is consistent for τ .
- ▶ This algorithm is known as the two-stage least squares (2SLS).

Two-stage least squares

- ▶ There are two other approaches to estimate τ .
- ▶ The first stage is always necessary.
- ▶ Let's denote the predicted value of D_i as \hat{D}_i and the regression residual as $\hat{\nu}_i$.
- ▶ The second approach is to regress Y_i on \hat{D}_i (the second stage).
- ▶ The OLS estimate will be consistent for τ .
- ▶ Consider the matrix form of the equations:

$$\mathbf{Y} = \mathbf{D}\tau + \varepsilon,$$

$$\mathbf{D} = \mathbf{Z}\delta + \nu.$$

Two-stage least squares

- Remember that

$$\hat{\delta} = (\mathbf{Z}'\mathbf{Z})^{-1}(\mathbf{Z}'\mathbf{D}),$$

$$\hat{\mathbf{D}} = \mathbf{Z}(\mathbf{Z}'\mathbf{Z})^{-1}(\mathbf{Z}'\mathbf{D}) = \mathbf{P}_Z\mathbf{D},$$

$$\hat{\nu} = \mathbf{D} - \hat{\mathbf{D}} = (\mathbf{I} - \mathbf{P}_Z)\mathbf{D} = \mathbf{Q}_Z\mathbf{D}.$$

- The OLS estimate from regressing \mathbf{Y} on $\hat{\mathbf{D}}$ equals:

$$\begin{aligned}\hat{\tau}_{2SLS} &= (\hat{\mathbf{D}}'\hat{\mathbf{D}})^{-1} (\hat{\mathbf{D}}'\mathbf{Y}) \\ &= (\mathbf{D}'\mathbf{P}_Z\mathbf{P}_Z\mathbf{D})^{-1} (\mathbf{D}'\mathbf{P}_Z\mathbf{Y}) \\ &= (\mathbf{D}'\mathbf{P}_Z\mathbf{D})^{-1} (\mathbf{D}'\mathbf{P}_Z(\mathbf{D}\tau + \varepsilon)) \\ &= \tau + (\mathbf{D}'\mathbf{P}_Z\mathbf{D})^{-1} (\mathbf{D}'\mathbf{P}_Z\varepsilon) \rightarrow \tau.\end{aligned}$$

Two-stage least squares

- ▶ It is natural to see that

$$\text{Var}[\hat{\tau}_{2SLS}] = (\mathbf{D}'\mathbf{P}_Z\mathbf{D})^{-1} (\mathbf{D}'\mathbf{P}_Z\varepsilon\varepsilon'\mathbf{P}_Z\mathbf{D}) (\mathbf{D}'\mathbf{P}_Z\mathbf{D})^{-1}.$$

- ▶ It is more complicated than the variance for the OLS estimator but also takes the sandwich form.
- ▶ Intuitively, we first project D_i onto the space spanned by Z_i .
- ▶ Next, we project Y_i onto the same space, with the projected D_i as the bases.
- ▶ We can estimate the variance by replacing ε with the regression residuals in the second stage.
- ▶ Incorporating covariates is straightforward by replacing \mathbf{D} with $\tilde{\mathbf{D}} = (\mathbf{D}, \mathbf{X})$ and \mathbf{Z} with $\tilde{\mathbf{Z}} = (\mathbf{Z}, \mathbf{X})$.

Control function

- ▶ The third approach is known as “control function” in econometrics.
- ▶ In the second stage, we regress Y_i on D_i and $\hat{\nu}_i$.
- ▶ We can show that the estimated coefficient for D_i will be consistent for τ .
- ▶ Note that ν_i is the endogenous part in D_i and $\hat{\nu}_i$ is unbiased for ν_i .
- ▶ If we can control for the endogenous part, the estimation will be unbiased.
- ▶ These approaches give you numerically equivalent results when the models are correct.

Control function

- ▶ Define the residual-making matrix with regards to $\hat{\nu}$:

$$\mathbf{Q}_{\hat{\nu}} = \mathbf{I} - \hat{\nu}(\hat{\nu}'\hat{\nu})^{-1}\hat{\nu}' = \mathbf{I} - \mathbf{Q}_Z\mathbf{D}(\mathbf{D}'\mathbf{Q}_Z\mathbf{D})^{-1}\mathbf{D}'\mathbf{Q}_Z.$$

- ▶ We can see that $\mathbf{Q}_{\hat{\nu}}\mathbf{D} = \mathbf{D} - \mathbf{Q}_Z\mathbf{D} = \mathbf{P}_Z\mathbf{D}$.
- ▶ Multiplying $\mathbf{Q}_{\hat{\nu}}$ on both sides of the main equation, we have

$$\tilde{\mathbf{Y}} = \mathbf{Q}_{\hat{\nu}}\mathbf{Y} = \mathbf{Q}_{\hat{\nu}}\mathbf{D}\tau + \mathbf{Q}_{\hat{\nu}}\varepsilon = \tilde{\mathbf{D}} + \tilde{\varepsilon}.$$

- ▶ From the FWL theorem,

$$\begin{aligned}\hat{\tau}_{CF} &= (\tilde{\mathbf{D}}'\tilde{\mathbf{D}})^{-1}(\tilde{\mathbf{D}}'\tilde{\mathbf{Y}}) = \tau + (\tilde{\mathbf{D}}'\tilde{\mathbf{D}})^{-1}(\tilde{\mathbf{D}}'\tilde{\varepsilon}) \\ &= \tau + (\mathbf{D}'\mathbf{Q}_{\hat{\nu}}\mathbf{D})^{-1}(\mathbf{D}'\mathbf{Q}_{\hat{\nu}}\varepsilon) \\ &= \tau + (\mathbf{D}'\mathbf{P}_Z\mathbf{D})^{-1}(\mathbf{D}'\mathbf{P}_Z\varepsilon) = \hat{\tau}_{2SLS}.\end{aligned}$$

Control function: application

The OLS estimate is 3.90903

The 2SLS estimate is 2.905916

The 2SLS estimate is 2.905916

The control function estimate is 2.905916

From 2SLS to LATE

- ▶ We have reviewed the literature on instrumental variables following the econometric tradition.
- ▶ We didn't not use the potential outcome notations at all!
- ▶ Are the 2SLS estimates causal from the design-based perspective?
- ▶ Note that the triangular system satisfies Assumptions 1-4 for identifying the LATE.
- ▶ Random assignment results from that $Z_i \perp \nu_i$.
- ▶ Exclusion restriction and the first stage hold due to the functional form.
- ▶ Monotonicity is assumed for $g()$.

From 2SLS to LATE

- ▶ Angrist, Imbens, and Rubin (1996) first showed that when D_i and Z_i are binary, $\hat{\tau}_{2SLS} = \hat{\tau}_{Wald}$.
- ▶ The result holds even when the effects are heterogeneous.
- ▶ In other words, $\hat{\tau}_{2SLS}$ can be interpreted as an estimate for the average treatment effect on the compliers (LATE).
- ▶ Compliers are the agents who are encouraged to select into the treatment by the instrument.
- ▶ We start from an economic model and end up with an interpretation rooted in the potential outcomes framework.
- ▶ The economic approach and the statistical approach are capturing the same concepts.
- ▶ We can justify the potential outcomes framework with social science theory, or analyze social science problems with the idea of counterfactual.
- ▶ This finding won Angrist and Imbens a Nobel prize.

From 2SLS to LATE

- ▶ In this case, the triangular system boils down to:

$$\begin{aligned}Y_i &= \tau_i D_i + \varepsilon_i, \\D_i &= \mathbf{1}\{(\delta_i Z_i + \nu_i) \geq 0\}.\end{aligned}$$

- ▶ We can see that $D_i(0) = \mathbf{1}\{\nu_i \geq 0\}$ and $D_i(1) = \mathbf{1}\{\delta_i + \nu_i \geq 0\}$.
- ▶ δ_i and ν_i determine the response of unit i to Z_i hence i 's type.
- ▶ For example, compliers are those with $-\delta_i \leq \nu_i < 0$.
- ▶ The analysis in Angrist, Imbens, and Rubin (1996) indicates that the 2SLS estimator's result is completely driven by these units.

2SLS: application

The LATE is 1.855

Estimate from the Wald estimator is 0.107

The 2SLS estimate is 0.107

SE of the 2SLS estimate is 0.914

From 2SLS to LATE

- ▶ These results led to the credibility revolution in the 90s.
- ▶ Using IVs in observational studies to identify the LATE became a fad.
- ▶ Angrist (1990): draft lottery in the Vietnam war - veteran status - earnings on the labor market.
- ▶ Angrist and Krueger (1991): birth season - age when dropping out of high school - earnings on the labor market.
- ▶ Acemoglu, Johnson, and Robinson (2001): settlers' mortality - inclusive institutions - economic development.
- ▶ But many applications turned out to be not as credible as we thought.

References I

- Acemoglu, Daron, Simon Johnson, and James A Robinson. 2001. "The Colonial Origins of Comparative Development: An Empirical Investigation." *American Economic Review* 91 (5): 1369–1401.
- Angrist, Joshua D. 1990. "Lifetime Earnings and the Vietnam Era Draft Lottery: Evidence from Social Security Administrative Records." *The American Economic Review*, 313–36.
- Angrist, Joshua D, Guido W Imbens, and Donald B Rubin. 1996. "Identification of Causal Effects Using Instrumental Variables." *Journal of the American Statistical Association* 91 (434): 444–55.
- Angrist, Joshua D, and Alan B Krueger. 1991. "Does Compulsory School Attendance Affect Schooling and Earnings?" *The Quarterly Journal of Economics* 106 (4): 979–1014.